# Lecture 14 Bayesian Models for Spatio-Temporal Data

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# Outline of Lecture

- 1 Recap of Bayesian Models
- 2 Empirical Bayes
- **3** Case 1: Long-Lead Forecasting of Sea Surface Temperatures
- 4 Case 2: Modeling and Forecasting the Eurasian Dove Invasion
- **5** Case 3: Mediterranean Surface Vector Winds
- 6 Wrapping Up the Course

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# **Bayesian Models**

- Bayesian models differ from frequentist models only in that the parameters  $\theta$  are random.
- This allows us to stack priors to create hierarchical models.
- The Gibbs sampler is a universal algorithm that allows us to efficiently sample from the posterior in hierarchical models.

## Example: Rater Model



# Computations

- The Gibbs sampler provides samples from the posterior.
- We can use these samples to estimate the posterior distribution (e.g., histogram) or the posterior mean.
- JAGS will automatically simulate from the posterior.

# Results for the Rater Model

I put priors on  $\gamma_k$  and  $\delta_\ell$  so that I could do inference on them.



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## What would a frequentist do?

- **Remember:** In the Bayesian framework, to perform inference on a parameter, you must put a prior on it. Otherwise, you must specify its value beforehand.
- **Example:** Suppose there are 3 sections of a class taught by different professors. Let  $y_{ij}$  denote the final exam score of student j in class i.
  - $y_{ij} | \theta_i \sim N(\theta_i, \sigma^2)$  (instructor effect)
  - $\theta_i \sim N(\mu, \tau^2)$
- Maybe we can instead try to estimate hyperparameters such as  $\mu$  and  $\tau^2$  from the data, then use  $N(\hat{\mu},\hat{\tau}^2)$  as the prior.
- What would our estimates of the parameters be then?

## **Empirical Bayes**

- The idea of estimating hyperparameters from the data is called *empirical Bayes*.
- It is ultimately a frequentist method because we don't need to specify a prior on the hyperparameters we estimate!
- We get hierarchical models without subjective priors! Is this too good to be true?
- What are the challenges of doing this in practice?

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# Problem Setup

- El Niño is characterized by warmer sea surface temperatures (SST) in the equatorial Pacific Ocean.
- Therefore, to predict El Niño, one needs to forecast the SST several months in advance.
- We observe the average monthly SST at different locations in the Pacific:

$$oldsymbol{z}_t \stackrel{def}{=} (oldsymbol{z}_t(oldsymbol{s}_1),...,oldsymbol{z}_t(oldsymbol{s}_m))$$

- Want to forecast SSTs au months in advance:  $z_{T+ au}$ .
- This analysis is taken from Berliner, Wikle, and Cressie (2000).

## First Model: Linear Dynamical Model

$$\begin{array}{lll} \text{State model:} & \boldsymbol{y}_{t+\tau} = \Phi \boldsymbol{y}_t + \boldsymbol{\epsilon}_t & \boldsymbol{\epsilon}_t \sim N(0, \Sigma) \\ \text{Data model:} & \boldsymbol{z}_t = A \boldsymbol{y}_t + \boldsymbol{\delta}_t & \boldsymbol{\delta}_t \sim N(0, \sigma^2 I) \end{array}$$

- A contains the first k principal components of the empirical covariance matrix over the spatial locations.
- $y_t$  represents weights on those PCs.
- Unknown parameters are  $\Phi, \Sigma, \sigma^2$ . Need to put priors on all of these:

$$\begin{split} & \mathsf{vec}(\Phi) \sim N(\mathsf{vec}(0.9I), 100I) \\ & \Sigma^{-1} \sim \mathsf{Wishart}\left(\frac{1}{100(k-1)}I, k-1\right) \\ & \sigma^2 \sim \mathsf{InverseGaussian}(0.1, 100) \end{split}$$

Case 1: Long-Lead Forecasting of Sea Surface Temperatures

### First Model: Linear Dynamical Model



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## Second Model: "Non-Linear" Model

- Allow  $\Phi_t = \Phi(I_t, J_t)$  to vary with time.
- *I<sub>t</sub>* classifies the current regime as "cool", "normal", or "warm". Obtained by thresholding the Southern Oscillation Index (SOI):

$$I_t = \begin{cases} 0 & \text{if } SOI_t < \text{low threshold} \\ 1 & \text{if } SOI_t \text{ in between} \\ 2 & \text{if } SOI_t > \text{upper threshold} \end{cases}$$

•  $J_t$  is obtained by similarly thresholding a latent process  $W_t$ :

$$W_t | \boldsymbol{\beta}, \tau^2 \sim N(\boldsymbol{x}_t^T \boldsymbol{\beta}, \tau^2)$$

## Second Model: "Non-Linear" Model



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## Problem Setup

- The Eurasian Collared Dove (ECD) was first observed in North America in the 80s and are now spreading quickly throughout the continent.
- They pose a threat to native ecosystems, so we would like to forecast their spread.
- Observe  $z_t(s_i)$ : number of doves observed at location  $s_i$ .
- This analysis is taken from Hooten, Wikle, Dorazio, and Royle (2007).















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### The Model

Data model: 
$$z_t(s_i) \mid y_t(s_i), \pi \sim \text{Bin}(y_t(s_i), \pi)$$
  
 $y_t \mid \lambda_t \sim \text{Pois}(H\lambda_t)$   
State model:  $\lambda_t = B(\alpha)G(\lambda_{t-1}; \theta)\lambda_{t-1}$ 

- $\pi$  is the probability of observing an animal. Not estimable from this data alone, but the authors estimated it using data collected on a related species.
- G is a diagonal matrix that models growth, while B models dispersion.
- The authors go on to put priors on  $\alpha$  and  $\theta$ .

#### Results

#### Posterior means for years in sample



#### Results

#### Posterior means for future years



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## Problem Setup

- The goal is to predict wind speeds and directions at different locations over the Mediterranean.
- $(x_t(s_i), y_t(s_i))$  is the vector indicating the wind speed in the *x* and *y*-directions.
- We assume that  $x_t$  and  $y_t$  are noisy measurements of underlying states  $u_t$  and  $v_t$ .

## Data on February 1, 2005



## A Physical State Model

• The state model is movtivated by the Rayleigh friction equations:

$$\frac{\partial u}{\partial t} - fv = -\frac{1}{\rho_0} \frac{\partial p}{\partial x} - \gamma u \qquad \quad \frac{\partial v}{\partial t} + fu = -\frac{1}{\rho_0} \frac{\partial p}{\partial y} - \gamma v.$$

where u, v are the east-west and north-south components, p is the sea-level pressure, and the rest are (unknown) parameters.

• An approximate solution to these equations is given by

$$\begin{split} u &\approx -\frac{f}{\rho_0(f^2+\gamma^2)}\frac{\partial p}{\partial y} - \frac{\gamma}{\rho_0(f^2+\gamma^2)}\frac{\partial p}{\partial x} - 2\frac{\gamma}{\rho_0(f^2+\gamma^2)}\frac{\partial u}{\partial t} - \frac{1}{\rho_0(f^2+\gamma^2)}\frac{\partial^2 p}{\partial x\partial t} \\ v &\approx \frac{f}{\rho_0(f^2+\gamma^2)}\frac{\partial p}{\partial x} - \frac{\gamma}{\rho_0(f^2+\gamma^2)}\frac{\partial p}{\partial y} - 2\frac{\gamma}{\rho_0(f^2+\gamma^2)}\frac{\partial v}{\partial t} - \frac{1}{\rho_0(f^2+\gamma^2)}\frac{\partial^2 p}{\partial y\partial t} \end{split}$$

• We can discretize this as follows:

$$u_t = a_1 D_y p_t + a_2 D_x p_t + a_3 u_{t-1} + a_4 D_x p_{t-1}$$
$$v_t = b_1 D_x p_t + b_2 D_y p_t + b_3 v_{t-1} + b_4 D_y p_{t-1}$$

### Results for February 2, 2005



BHM, 10 realizations and post. mean, 2/2/2005 18:00, itime = 34, MedBhm\_real\_A2E3100K\_uv\_fmt

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# A Few Last Thoughts

- The analysis of spatial and temporal data is really the analysis of correlated data.
- "Model the mean function, use spatial and temporal methods to model the residual."
- There are two main ways to capture correlations: model the correlation directly (e.g., kriging) and via autoregressions.
- Many methods that work well for time series (e.g., Kalman filter) break down in space because the data are no longer ordered.

#### Thanks for a great quarter!